

BOSH TO'PLAMNING MATEMATIK KUTILISHINI TURLI HAJMLI TANLANMALAR BO'YICHA BAHOLASH UCHUN ISHONCHLI INTERVALLAR

Raximova Umida Ziyadullayevna

SamISI assistenti

ARTICLE INFO.	Annotatsiya
<p>Kalit so'zlar: Bosh to'plam, dispersiya, normal taqsimot, St'yudent taqsimoti, ishonchli interval.</p>	<p>Ushbu ish ilmiy va nazariy xulosalar hosil qilish maqsadida kuzatish natijasida to'plangan ma'lumotlarga asoslanib, taqsimotning noma'lum parametrini baholash usullarini ko'rsatib berishga bog'ishlangan.</p>

<http://www.gospodarkainnowacje.pl/> © 2023 LWAB.

Normal taqsimlangan bosh to'plamdan erkli ikkita n_1 va n_2 hajmli tanlanma olingan. \bar{X}_1 va \bar{X}_2 - bu tanlanmalar bo'yicha tanlanma o'rtacha qiymatlar, S_1^2 va S_2^2 esa tanlanma dispersiyalar bo'lsin. $n_1 + n_2$ lar hajmli birlashtirilgan tanlanma bo'yicha tanlanma o'rtacha qiymat va tanlanma dispersiya

$$\bar{X} = \frac{n_1 \bar{X}_1 + n_2 \bar{X}_2}{n_1 + n_2}, S^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

formular bo'yicha hisoblanadi.

Bosh to'plamning noma'lum μ matematik kutilishi uchun berilgan ishonchlilik ehtimolida ishonchli intervallar qurish talab etiladi.

Bu masalani yechishda ikki holni qarymiz.

1. Bosh to'plam dispersiyasi σ^2 ma'lum bo'lgan hol. X_1, X_2, \dots, X_{n_1} va Y_1, Y_2, \dots, Y_{n_2} lar shu bosh to'plamdan olingan n_1 va n_2 hajmli tanlanmalar bo'lsin.

Ushbu $Z = \frac{\bar{X} - \mu}{\sigma} \sqrt{n_1 + n_2}$ tasodifiy miqdorlarni qaraymiz. Agar bosh to'plam normal taqsimlangan bo'lsa, u holda erkli kuzatishlar bo'yicha topilgan

$$\bar{X}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} X_i, \bar{X}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} X_i,$$

tanlanma o'rtacha qiymatlar ham normal taqsimlangan bo'ladi. \bar{X}_1 taqsimotning parametrlari

$$M\bar{X}_1 = \mu, \sigma_{x_1} = \sqrt{D\bar{X}_1} = \frac{\sigma}{\sqrt{n_1}}, \bar{X}_2 \text{ taqsimotning parametrlari esa}$$

$$M\bar{X}_2 = \mu, \sigma_{x_2} = \sqrt{D\bar{X}_2} = \frac{\sigma}{\sqrt{n_2}} \text{ bo'ladi.}$$

Haqiqatdan ham $X_i(Y_i)$ qiymatlarning har biri va bosh to'plam belgisi bir xil taqsimotga ega ekanini e'tiborga olsak, ularning va bosh to'plam belgisining son harakteriskalari bir xil bo'ladi. Jumladan $X_i(Y_i)$ miqdorlarning har birining matematik kutilishi bosh to'plam belgisining matematik kutilishiga dispersiyasi esa bosh to'plam belgisining dispersiyasiga, ya'ni

$$MX_i = \mu, MY_i = \mu, Dx_i = \sigma^2, DY_i = \sigma^2.$$

U holda

$$\begin{aligned} M\bar{X}_1 &= M\left(\frac{1}{n_1} \sum_{i=1}^{n_1} X_i\right) = \frac{1}{n_1} M\left(\sum_{i=1}^{n_1} X_i\right) = \\ &= \frac{1}{n_1} \sum_{i=1}^{n_1} MX_i = \frac{1}{n_1} \cdot n_1 \mu = \mu, \\ \sigma_{x_1}^2 &= D\bar{X}_1 = D\left(\frac{1}{n_1} \sum_{i=1}^{n_1} X_i\right) = \frac{1}{n_1^2} D\left(\sum_{i=1}^{n_1} X_i\right) = \\ &= \frac{1}{n_1^2} \sum_{i=1}^{n_1} DX_i = \frac{1}{n_1^2} \cdot n_1 \sigma^2 = \frac{\sigma^2}{n_1}, \quad \sigma_{x_1} = \frac{\sigma}{\sqrt{n_1}} \end{aligned}$$

Huddi shunday

$$\begin{aligned} M\bar{X}_2 &= M\left(\frac{1}{n_2} \sum_{i=1}^{n_2} Y_i\right) = \frac{1}{n_2} M\left(\sum_{i=1}^{n_2} Y_i\right) = \\ &= \frac{1}{n_2} \sum_{i=1}^{n_2} MY_i = \frac{1}{n_2} \cdot n_2 \mu = \mu, \\ \sigma_{x_2}^2 &= D\bar{X}_2 = D\left(\frac{1}{n_2} \sum_{i=1}^{n_2} Y_i\right) = \frac{1}{n_2^2} D\left(\sum_{i=1}^{n_2} Y_i\right) = \\ &= \frac{1}{n_2^2} \sum_{i=1}^{n_2} DY_i = \frac{1}{n_2^2} \cdot n_2 \sigma^2 = \frac{\sigma^2}{n_2}, \quad \sigma_{x_2} = \frac{\sigma}{\sqrt{n_2}} \end{aligned}$$

Shuningdek, \bar{X} tasodifiy miqdor ham normal taqsimlangan bo'lib, \bar{X} taqsimotning parametrlari

$$M\bar{X} = \mu, D\bar{X} = \frac{\sigma^2}{n_1 + n_2} \text{ bo'ladi.}$$

Haqiqatdan

$$\begin{aligned}
 M\bar{X} &= M\left(\frac{n_1\bar{X}_1 + n_2\bar{X}_2}{n_1 + n_2}\right) = \frac{1}{n_1 + n_2} M(n_1\bar{X}_1 + n_2\bar{X}_2) = \\
 &= \frac{1}{n_1 + n_2} [M(n_1\bar{X}_1) + M(n_2\bar{X}_2)] = \frac{1}{n_1 + n_2} (n_1M\bar{X}_1 + n_2M\bar{X}_2) = \\
 &= \frac{1}{n_1 + n_2} (n_1\mu + n_2\mu) = \mu, \\
 D\bar{X} &= D\left(\frac{n_1\bar{X}_1 + n_2\bar{X}_2}{n_1 + n_2}\right) = \frac{1}{(n_1 + n_2)^2} D(n_1\bar{X}_1 + n_2\bar{X}_2) = \\
 &= \frac{1}{(n_1 + n_2)^2} [D(n_1\bar{X}_1) + D(n_2\bar{X}_2)] = \frac{1}{(n_1 + n_2)^2} (n_1^2 D\bar{X}_1 + n_2^2 D\bar{X}_2) = \\
 &= \frac{1}{(n_1 + n_2)^2} \left(n_1^2 \cdot \frac{\sigma^2}{n_1} + n_2^2 \cdot \frac{\sigma^2}{n_2}\right) = \frac{\sigma^2}{n_1 + n_2}
 \end{aligned}$$

Z tasodifiy miqdor ham \bar{X} normal argumentning chiziqli funksiyasi sifatida normal taqsimotga ega bo'lib, taqsimotning parametrlari $MZ = 0$, $DZ = 1$ bo'ladi.

Haqiqatdan

$$\begin{aligned}
 MZ &= M\left(\frac{\bar{X} - \mu}{\sigma} \sqrt{n_1 + n_2}\right) = \frac{\sqrt{n_1 + n_2}}{\sigma} M(\bar{X} - \mu) = \\
 &= \frac{\sqrt{n_1 + n_2}}{\sigma} (M\bar{X} - \mu) = \frac{\sqrt{n_1 + n_2}}{\sigma} (\mu - \mu) = 0, \\
 DZ &= D\left(\frac{\bar{X} - \mu}{\sigma} \sqrt{n_1 + n_2}\right) = \frac{n_1 + n_2}{\sigma^2} D(\bar{X} - \mu) = \\
 &= \frac{n_1 + n_2}{\sigma^2} (D\bar{X} - D\mu) = \frac{n_1 + n_2}{\sigma^2} (D\bar{X} - 0) = \\
 &= \frac{n_1 + n_2}{\sigma^2} \cdot \frac{\sigma^2}{n_1 + n_2} = 1
 \end{aligned}$$

Undan tashqari Z tasodifiy miqdor μ ning funksiyasi sifatida uzluksiz va manoton hamdir. Ushbu

$$P(U_{\alpha/2} < Z < U_{1-\alpha/2}) = P\left(U_{\alpha/2} < \frac{\bar{X} - \mu}{\sigma} \sqrt{n_1 + n_2} < U_{1-\alpha/2}\right) = 1 - \alpha$$

munosabat bajarilishini talab qilamiz, bu yerda $U_{\alpha/2}$ va $U_{1-\alpha/2}$ - matematik kutilishi 0 ga, dispersiyasi esa 1 ga teng bo'lgan normal taqsimotning kvantillari.

Qavs ichidagi

$$U_{\alpha/2} < \frac{\bar{X} - \mu}{\sigma} \sqrt{n_1 + n_2} < U_{1-\alpha/2}$$

tengsizlikni μ ga nisbatan yechib, $1 - \alpha$ ehtimol bilan

$$\bar{X} - \frac{\sigma}{\sqrt{n_1 + n_2}} U_{1-\alpha/2} < \mu < \bar{X} - \frac{\sigma}{\sqrt{n_1 + n_2}} U_{\alpha/2}$$

tengsizlik bajarilishini hosil qilamiz.

Normal taqsimotning kvantillari $U_{\alpha/2} = -U_{1-\alpha/2}$ bog'lanishda bo'lganidan, μ uchun hosil qilingan ishonchli intervalni quyidagicha yozish mumkin:

$$\bar{X} - \frac{\sigma}{\sqrt{n_1 + n_2}} U_{1-\alpha/2} < \mu < \bar{X} + \frac{\sigma}{\sqrt{n_1 + n_2}} U_{1-\alpha/2}.$$

2. Bosh to'plam dispersiyasi σ^2 noma'lum bo'lgan hol. Ravshanki, bu holda oldingi band natijalaridan foydalanib bo'lmaydi, chunki u yerda σ^2 ma'lum deb faraz qilinga edi. Hozirgi holda markaziy statistika sifatida

$$T = \frac{\bar{X} - \mu}{S} \sqrt{n_1 + n_2}$$

tasodifiy miqdor qabul qilinadi. T tasodifiy miqdor $k = n_1 + n_2 - 2$ ozodlik darajali St'yudent taqsimotiga ega bo'ladi. St'yudent taqsimoti $n_1 + n_2$ parametr bilan aniqlanishini, μ va σ parametrlarga bog'liq emasligini osongina ko'rish mumkin. Demak,

$$\begin{aligned} P\{t_{\alpha/2}(n_1 + n_2 - 2) < T < t_{1-\alpha/2}(n_1 + n_2 - 2)\} = \\ = P\{t_{\alpha/2}(n_1 + n_2 - 2) < \frac{\bar{X} - \mu}{S} \sqrt{n_1 + n_2} < t_{1-\alpha/2}(n_1 + n_2 - 2)\} = 1 - \alpha \end{aligned}$$

bu yerda $t_{\alpha/2}(n_1 + n_2 - 2)$ va $t_{1-\alpha/2}(n_1 + n_2 - 2)$ lar $k = n_1 + n_2 - 2$ ozodlik darajali St' yudent taqsimotning kvantillari.

Qavis ichidagi

$$t_{\alpha/2}(n_1 + n_2 - 2) < \frac{\bar{X} - \mu}{S} \sqrt{n_1 + n_2} < t_{1-\alpha/2}(n_1 + n_2 - 2)$$

tengsizlikni μ ga nisbatan yechib,

$$\bar{X} - \frac{S}{\sqrt{n_1 + n_2}} t_{1-\alpha/2}(n_1 + n_2 - 2) < \mu < \bar{X} - \frac{S}{\sqrt{n_1 + n_2}} t_{\alpha/2}(n_1 + n_2 - 2)$$

ni hosil qilamiz.

St'yudent taqsimotining kvantillari $t_{\alpha/2}(n_1 + n_2 - 2) = -t_{1-\alpha/2}(n_1 + n_2 - 2)$ day bog'lanishdaligini hisobga olib, berilgan $1 - \alpha$ ishonchlilik ehtimolida μ uchun ishonchli intervalni ,ya'ni

$$\bar{X} - \frac{S}{\sqrt{n_1 + n_2}} t_{1-\alpha/2}(n_1 + n_2 - 2) < \mu < \bar{X} + \frac{S}{\sqrt{n_1 + n_2}} t_{1-\alpha/2}(n_1 + n_2 - 2)$$

ni hosil qilamiz.

Adabiyotlar

1. Э.А. Вукулов и др. Теория вероятностей и математическая статистика. М., Наука, 1990, 428 с.
2. Г.И. Ивченко, Ю.И. Медведев. Математическая статистика. М., Высшая школа, 1984., 248 с.